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Rapid Protective Effects of Early BCG on Neonatal Mortality Among Low Birth Weight Boys: Observations From Randomized Trials

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Background. Three randomized trials (RCTs) in low-weight (<2.5 kg) infants have shown that Bacille Calmette-Guérin (BCG) vaccine nonspecifically reduces all-cause mortality in the neonatal period.

Methods. Using data from 3 RCTs of early BCG (n = 6583) we examined potential sex differences in the timing of the mortality reduction in the neonatal period, presenting metaestimates of the main outcome mortality rate ratios (MRR) for BCG-vaccinated and controls.

Results. Among controls, boys had a particularly high mortality during the first week after randomization: male–female MRR 2.71 (95% CI, 1.70–4.50). During the first week, BCG had a marked beneficial effect for boys, reducing mortality 3-fold (MRR [BCG/no BCG] = 0.36 [0.20–0.67]). In weeks 2–4 the effect waned for boys (MRR = 0.91 [0.51–1.69]). In girls, the pattern was opposite with a limited effect in the first week (MRR = 0.85 [0.46–1.54]), but a significant reduction in weeks 2–4 (MRR = 0.56 [0.31–1.00]). This was consistent in all 3 trials. Verbal autopsies linked early benefit to fewer sepsis-related deaths among BCG-vaccinated boys.

Discussion. The marked reduction in mortality in the days after BCG vaccination in boys emphasizes the importance of providing BCG soon after birth.

Trial registration numbers: ClinicalTrials.gov (NCT00146302) and ClinicalTrials.gov (NCT00625482).

Keywords: Bacille Calmette-Guérin vaccine; heterologous immunity; nonspecific effects of vaccines; sex differences; neonatal mortality.

Observational studies have suggested that Bacille Calmette-Guérin (BCG) vaccine is associated with lower all-cause mortality in addition to protecting against tuberculosis [1, 2]. We have called this phenomenon the nonspecific effects of BCG.

In many low-income countries, including Guinea-Bissau, BCG is postponed for low-birth-weight (<2500 g) neonates until they have attained a weight of 2500 g. It was therefore possible to conduct randomized trials (RCTs) among low-weight (LW) infants to test the hypothesis that BCG at birth reduced infant mortality by 25%. Between 2002 and 2008, we conducted 3 RCTs of early BCG showing a nonsignificant reduction of 21% in infant mortality, but a significant reduction of 48% in neonatal mortality [3]. Because the first 2 trials had infant mortality as the primary outcome, we subsequently conducted a third RCT (2008–2013) with neonatal mortality as the main outcome. A combined analysis of all 3 trials showed significant reductions of 38% in neonatal mortality and 16% in infant mortality [4]. The 3 trials supported consistently that BCG has strong beneficial effects on mortality in the first month of life.

While analyzing the first 2 RCTs we observed that the beneficial effects of BCG had different timing for boys and girls. Before publishing, we awaited the results of the third trial, which has now been completed [4]. Here, we present the sex-differential time trend in the effect of BCG on neonatal mortality in the 3 RCTs conducted in LW infants in Guinea-Bissau.

METHODS

Background

The Bandim Health Project (BHP) maintains a health and demographic surveillance system in Bissau, the capital of Guinea-Bissau. The first LW trial began in November 2002 (ClinicalTrials.gov, NCT00146302) [3]. All LW children from the city of Bissau born at the maternity ward of the national hospital, as well as LW children coming for vaccination to the health centers in the study area, were invited to participate in the trial. In November 2004, we discovered that there were faulty randomization procedures at the main hospital and the trial was stopped. The randomization procedures occurred without error at the health centers (Trial...
From May 2005 to January 2008 (Trial 2), we also studied the effect of vitamin A supplementation (VAS) at birth in a 2-by-2 factorial design. The children were randomly assigned to early BCG or delayed BCG as well as vitamin A supplementation or placebo. Because there was no interaction in mortality between the 2 interventions [3], the effects of BCG and VAS were reported separately [6].

Children living outside the BHP study area were driven home by a field team. The team drew a map of the house, recorded global positioning system (GPS) coordinates, and took a photograph of the house and the mother to ensure that the team would be able to identify the child at subsequent visits. The children were visited at home after 3 days and at 2, 6, and 12 months of age to assess the status of the child. Children from the study area were identified via their unique study number and followed both by the routine visits by the BHP assistants and at the study specific visits at 3 days and at 2, 6, and 12 months of age.

**Verbal Autopsy Procedures**

Verbal autopsies were conducted to determine probable causes of death. During Trial 1 and 2, a medical practitioner visited the home 3 months after the child had died. The medical practitioner registered the events leading up to the death of the child, and assigned a cause of death based on this information. In Trial 3, a field assistant trained in performing verbal autopsies visited the home 3 months after the child had died. The field assistant recorded events leading up to the death of the child. Based on this information, the medical practitioner, who had assigned the causes of death in Trial 1 and 2, assigned the causes of death in Trial 3. The medical practitioner was unaware of the intervention allocation.

**Statistical Analysis**

Cox proportional hazards models with time since randomization as the underlying time variable were used to compare mortality rates between the intervention and the control group, providing mortality rate ratios (MRR). Children entered the analyses at the day of randomization and exited at death or at the end of the neonatal period (28 days of age). Children who moved before 28 days of age were censored. We used robust 95% confidence intervals (CI) to adjust the standard errors for the lifetime dependence between twins. Time from randomization to 28 days was divided into intervals, and within these intervals we analyzed the effect of early BCG. Because most deaths occurred in the first week after randomization, we divided it into 2 intervals of 0–3 days and 4–7 days after randomization; the remaining time of the neonatal period was analyzed in weekly intervals. We allowed a different baseline hazard for each of the 3 trials. When there were 0 deaths in the intervention group (MRR = 0), CIs were obtained by a profile likelihood method, inverting...
results of the likelihood ratio test into a CI. We evaluated the proportional-hazards assumption by log-log plots and Schoenfeld residuals. Kaplan-Meier curves with time since randomization as the underlying time were drawn to display the cumulative mortality between the intervention and control group for Trials 1–3 combined and for Trials 2 and 3 separately (Trial 1 had too few deaths for graphic presentation). All analyses were performed separately for boys and girls because the focus of the present study was to determine whether the effect of BCG differed by sex.

Ethics

The Guinean Ministry of Health’s Research Coordination Committee approved all trials and we received consultative approval from the Danish Central Ethical Committee for all trials. Children enrolled in the study had access to free consultations and essential drugs at the health centers in the BHP study area.

Participant Involvement

No participants were actively involved in setting the research question or the outcome measures, nor were they involved in the analysis, interpretation, and writing of the results. The participants are thanked in the acknowledgments. Our findings from the study have been shared with the Ministry of Health in Guinea-Bissau, to inform policy.

RESULTS

In Trial 1, 105 children were included and 1 child was excluded (Figure 1). In Trial 2, 2343 children were invited to participate and 23 were excluded. In Trial 3, 4193 children were invited and 34 were excluded. A total of 58 children were enrolled after the neonatal period and were not included in the present analysis. We therefore analyzed 6525 children and among them 189 deaths occurred during the first 28 days of life. At baseline, the randomization groups were comparable in each of the 3 trials [3–5]. In the combined population of the 3 trials, the intervention and control groups were comparable when stratified by sex (Supplementary Table 1). The median age of inclusion was 2 days (10–90 percentiles: 2–10 days) in all 3 trials.

For both sexes, the highest mortality rate was seen in the first 3 days after enrolment (Table 1, Figure 2). The mortality rate was much higher for control boys than for control girls particularly in the first 7 days after enrolment, the male-female MRR being 2.71 (95% CI, 1.64–4.50). After the first week, the MRR comparing boys and girls was reduced to 1.25 (95% CI, 0.70–2.22) (data not shown).

The mortality curves for the 2 randomization groups (early BCG and control) for all trials combined and individually for Trials 2 and 3 are depicted in Figure 3. The effect of BCG was consistently different for boys and girls for Trials 2 and 3 individually and for all 3 trials combined. Trial 1 had too few deaths to

![Figure 1](https://example.com/figure1.png)

**Figure 1.** Children enrolled in 3 trials of early BCG effects on neonatal mortality in Guinea-Bissau 2002–2013. Note: The flow of children in each trial has been described in detail elsewhere [3–5]. Abbreviation: BCG, Bacille Calmette-Guérin.
be presented graphically, although a strong early effect for boys was also observed in this trial (Table 2). Hence, the more rapidly occurring reduction in mortality after BCG for boys compared with girls was present in all 3 trials (Figure 3 and Table 2).

When analyzing all 3 trials combined, during the first 7 days after enrolment BCG versus control was associated with a significant reduction in mortality (1–3 days: MRR = 0.36 [95% CI, 0.16–0.80] and 4–7 days: MRR = 0.37 [95% CI, 0.14–0.94]) for boys (Table 1). In girls, there was only a very small effect of BCG at 1–3 days (MRR = 0.84 [95% CI, 0.41–1.73]) and 4–7 days (MRR = 0.85 [95% CI, 0.29–2.53]). This resulted in an interaction between sex and BCG in the first 7 days (P = .05 for interaction). In the following 3 weeks, boys had only a limited response to BCG (MRR = 0.91 [95% CI, 0.51–1.69]); however, girls had a significant 44% reduction in mortality from day 8 to 28 after birth (MRR: 0.56 [95% CI, 0.31–1.00]). Hence, the effect of BCG in boys and girls was significantly different in the first 1–7 days compared to the subsequent 8–28 days (P = .03 for 3-way interaction). We also tested other cut points from day 1 to 14 to verify that the 3-way interaction between BCG, sex, and time interval was not a chance observation (Supplementary Table 2). For all cut points in the first 14 days after enrolment, we observed the pattern of a stronger beneficial effect for boys in the first time interval and a stronger effect for girls in the second time interval. However, the strongest differential effect was seen for cut points between 4 and 10 days after enrolment and the most significant 3-way interaction was seen at 6 days after enrolment (P = .02 for interaction). When stratifying by trial, the same sex-differential time course effect of BCG was evident in all 3 trials (Table 2).

We had verbal autopsy information on cause of death from 95% (179/189) of the deaths that occurred in the neonatal period (Table 3). Sepsis was the cause of 50% of the deaths (90/179), other infections accounted for 19% (34/179) and noninfectious causes were 31% (55/179). For both sexes combined, BCG reduced sepsis-related mortality by 43% (MRR = 0.57 [95% CI, 0.36–0.91]) and other infections by 52% (MRR = 0.48 [95% CI, 0.21–1.08]), whereas the effect was limited for noninfectious causes (MRR = 0.89 [95% CI, 0.50–1.58]). Among boys, there was a significant reduction in sepsis-related deaths at 1–28 days (MRR = 0.51 [95% CI, 0.26–0.95]); however, this occurred most strongly in the first week (MRR = 0.37 [95% CI, 0.14–0.88]); for girls the reduction in sepsis-related deaths was strongest in weeks 2–4 after enrolment (MRR = 0.46 [95% CI, 0.14–1.29]).

**DISCUSSION**

Using data from 3 RCTs of early BCG, the present analysis found a rapidly occurring protective effect on mortality in boys; the effect was strongest for boys in the first week after receiving BCG, whereas there was little or no effect in the remainder of the neonatal period. This pattern differed significantly from that seen in girls, who had the strongest beneficial effect in weeks 2–4.

**Strength and Weaknesses**

The time course of the protective effect of BCG on mortality has not been examined previously. All 3 RCTs showed a strong
early beneficial effect of BCG in boys, a pattern different from that seen in girls.

The loss to follow-up was very low and only 0.6% of the children did not have complete follow-up in the neonatal period. Therefore, we do not believe that deaths were underreported in any group. The randomization was not blinded to the trial team, as the vaccinations were registered on the children’s health cards. This was done to ensure that control mothers would return to the health centers to get their children vaccinated with BCG when they had gained weight. The staff members working on the 3 trials were not involved in providing health care for study children, and health care providers at the hospital and health centers were not aware of the trials. It is therefore unlikely that lack of blinding affected subsequent treatment.
Table 2. Mortality Rate Ratios (MRR) in Time Intervals (0–7 days and 8–28 days) Since Randomization, Stratified by Sex and Trial

<table>
<thead>
<tr>
<th></th>
<th>Boys</th>
<th></th>
<th></th>
<th>Girls</th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Mortality per 100/PY (deaths/PY)</td>
<td>MRR for Early BCG versus CONTROL (95% CI)</td>
<td></td>
<td></td>
<td>Mortality per 100/PY (deaths/PY)</td>
<td>MRR for Early BCG versus CONTROL (95% CI)</td>
</tr>
<tr>
<td>Early BCG Group</td>
<td>Control Group</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
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<tr>
<td><strong>Trial 1</strong></td>
<td></td>
<td></td>
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<td></td>
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</tr>
<tr>
<td>1–7 days</td>
<td>0 (0/0.4)</td>
<td>1130 (4/0.4)</td>
<td><strong>0 (0–0.55)</strong></td>
<td>225 (1/0.4)</td>
<td>174 (1/0.6)</td>
<td>1.29 (0.08–20.28)</td>
</tr>
<tr>
<td>8–28 days</td>
<td>93 (1/10.7)</td>
<td>142 (1/7.1)</td>
<td>0.36 (0.02–7.00)</td>
<td>0 (0/9.6)</td>
<td>0 (0/12.3)</td>
<td>...</td>
</tr>
<tr>
<td><strong>Trial 2</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1–7 days</td>
<td>86 (8/9.3)</td>
<td>218 (20/9.2)</td>
<td><strong>0.40 (0.17–0.90)</strong></td>
<td>40 (5/12.6)</td>
<td>56 (7/12.4)</td>
<td>0.70 (0.22–2.22)</td>
</tr>
<tr>
<td>8–28 days</td>
<td>35 (8/21.6)</td>
<td>37 (7/21.6)</td>
<td>1.14 (0.38–3.91)</td>
<td>20 (6/30.2)</td>
<td>48 (14/29.5)</td>
<td>0.42 (0.16–1.09)</td>
</tr>
<tr>
<td><strong>Trial 3</strong></td>
<td></td>
<td></td>
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<td></td>
</tr>
<tr>
<td>1–7 days</td>
<td>48 (6/12.5)</td>
<td>113 (14/12.3)</td>
<td>0.42 (0.16–1.10)</td>
<td>53 (14/26.6)</td>
<td>60 (16/26.6)</td>
<td>0.84 (0.41–1.73)</td>
</tr>
<tr>
<td>8–28 days</td>
<td>41 (12/29.2)</td>
<td>49 (14/28.4)</td>
<td>0.83 (0.39–1.80)</td>
<td>19 (12/62.9)</td>
<td>28 (18/62.4)</td>
<td>0.68 (0.31–1.40)</td>
</tr>
</tbody>
</table>

Estimates with P < .05 are highlighted in bold.
Abbreviations: BCG, Bacille Calmette-Guérin; PY, person years.

Table 3. Mortality Rate Ratios (MRR) of Causes of Deaths Stratified by Sex, Presented as the Complete Period or Split into Time Intervals since Randomization

<table>
<thead>
<tr>
<th>Causes of Deaths</th>
<th>Boys</th>
<th></th>
<th></th>
<th>Girls</th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Mortality per 100/PY (deaths/PY)</td>
<td>MRR (BCG/Control)</td>
<td></td>
<td>Mortality per 100/PY (deaths/PY)</td>
<td>MRR (BCG/Control)</td>
<td></td>
</tr>
<tr>
<td></td>
<td>Early BCG</td>
<td>Control</td>
<td></td>
<td></td>
<td>Early BCG</td>
<td>Control</td>
</tr>
<tr>
<td>1–28 days after randomization(^a)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Sepsis</td>
<td>16/773</td>
<td>31/75.7</td>
<td><strong>0.50 (0.27–0.92)</strong></td>
<td>17/139.5</td>
<td>26/138.4</td>
<td>0.66 (0.36–1.21)</td>
</tr>
<tr>
<td>Other infections(^b)</td>
<td>7/773</td>
<td>14/75.7</td>
<td>0.49 (0.20–1.22)</td>
<td>4/139.5</td>
<td>10/138.4</td>
<td>0.40 (0.12–1.27)</td>
</tr>
<tr>
<td>Noninfections(^c)</td>
<td>11/773</td>
<td>13/75.7</td>
<td>0.82 (0.37–1.81)</td>
<td>15/139.5</td>
<td>15/138.4</td>
<td>1.00 (0.48–2.08)</td>
</tr>
<tr>
<td>Time intervals since randomization(^d)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Sepsis</td>
<td>8/22.2</td>
<td>21/21.9</td>
<td><strong>0.37 (0.17–0.84)</strong></td>
<td>11/171</td>
<td>13/171</td>
<td>0.86 (0.38–1.91)</td>
</tr>
<tr>
<td>Other infections(^b)</td>
<td>2/22.2</td>
<td>8/21.9</td>
<td>0.25 (0.05–1.16)</td>
<td>2/171</td>
<td>5/171</td>
<td>0.40 (0.08–2.06)</td>
</tr>
<tr>
<td>Noninfections(^c)</td>
<td>3/22.2</td>
<td>7/21.9</td>
<td>0.42 (0.11–1.64)</td>
<td>6/171</td>
<td>4/171</td>
<td>1.51 (0.43–5.37)</td>
</tr>
<tr>
<td>8–28 days after randomization</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Sepsis</td>
<td>8/55.1</td>
<td>10/53.8</td>
<td>0.76 (0.29–1.96)</td>
<td>6/99.7</td>
<td>13/98.7</td>
<td>0.46 (0.17–1.20)</td>
</tr>
<tr>
<td>Other infections(^b)</td>
<td>5/55.1</td>
<td>6/53.8</td>
<td>0.82 (0.25–2.68)</td>
<td>2/99.7</td>
<td>5/98.7</td>
<td>0.39 (0.06–2.03)</td>
</tr>
<tr>
<td>Noninfections(^c)</td>
<td>8/55.1</td>
<td>6/53.8</td>
<td>1.27 (0.44–3.62)</td>
<td>9/99.7</td>
<td>11/98.7</td>
<td>0.81 (0.32–2.02)</td>
</tr>
</tbody>
</table>

Estimates with P < .05 are highlighted in bold.
Abbreviations: BCG, Bacille Calmette-Guérin; PY, person years.
\(^a\)10 verbal autopsies were not conducted due to the respective families having moved outside the catchment area before the verbal autopsy visit.
\(^b\)The category ‘Other infections’ consists of deaths due to fever, respiratory infections, gastrointestinal infections, infections in the umbilical cord, malaria, or HIV.
\(^c\)The category ‘Noninfections’ consists of deaths due to prematurity, anemia, congenital problems, kernicterus, sudden infant death syndrome, gastrointestinal hemorrhage, convulsions, or bleeding.

Biological Mechanisms

The rapid protective effect of BCG vaccination on all-cause mortality in humans was first described in connection with the present trials [3–5]. A study in adult mice found that BCG administered 24 hour prior to challenge with *Klebsiella pneumoniae* significantly improved survival, and a protective effect was also seen when BCG was administered 1 hour after the challenge [7]. The rapid onset of these effects suggests that innate immune mechanisms are involved. The innate immune system can be stimulated nonspecifically, resulting in increased protection against unrelated infections, a phenomenon that has been called trained innate immunity [8]. Innate training effects of BCG have recently been demonstrated in healthy adult volunteers, in whom BCG was associated with increased proinflammatory in vitro cytokine responses to unrelated pathogens 2 weeks and 3 months after vaccination. Mechanistic studies showed that this was due to epigenetic reprogramming of monocytes [9]. In severe combined immune deficiency (SCID) mice, lacking an adaptive immune system, BCG was associated with no mortality after challenge with a *Candida* infection which was otherwise lethal in nonvaccinated mice [9]. These effects were not analyzed by sex. We have recently shown, in an immunological study of a subgroup of the infants enrolled in Trial 3, that BCG was associated with increased proinflammatory in vitro cytokine responses to innate agonists 4 weeks after vaccination [10], consistent with the findings from the
above mechanistic study in adults. Analyzed by sex, the effect estimates were larger in girls than in boys for most cytokine responses. The stronger immunological effect in girls 4 weeks after early BCG is consistent with the findings from the present study that 2–4 weeks after BCG vaccination there is a greater beneficial effect in females than males.

In the first week after vaccination, BCG decreased the high mortality in boys to a level comparable to that of girls. The sex-differential and time-dependent effects were particularly evident for sepsis-related deaths: BCG decreased sepsis-related mortality in boys within the first week after vaccination, with no immediate effect in girls, whereas a protective effect against sepsis-related mortality in girls was seen after the first week. It is well known that newborn boys experience infection-related disease and death more frequently than girls [11, 12]. The dissimilarity may arise from inherent differences in the immune function between boys and girls. However, very few studies have compared immune function in newborn boys and girls. High levels of interleukin-1 (IL-1) receptor antagonist in umbilical cord blood from preterm infants were associated with adverse outcome in girls but not in boys [13]. After uncomplicated pregnancies, cord blood responses to lipopolysaccharide stimulation were higher in boys than in girls for IL-1β and IL-6, but not tumor necrosis factor-alpha (TNF-α) [14]. The proportion of CD4 T cells in cord blood is higher and the CD4 and CD8 T cells express more Fas in girls compared with boys [15], indicative of lower apoptotic potential in boys. Although such sex-differences in early life immune status have been reported, we do not know whether they are related to the sex-differential timing of the beneficial effect of BCG.

Implications

Even though the World Health Organization (WHO) recommends BCG at birth to all infants, BCG is often administrated much later. Less than 50% receive BCG during the neonatal period in Africa [16]. There is no focus on delivering BCG very early and there is an inherent flaw in the health infrastructure to provide BCG early; for example, many maternity wards do not provide BCG before discharge and though WHO recommend home visits to newborn children these visits do not include vaccinations. Furthermore, there has been a counterproductive emphasis on reducing wastage of vaccine doses; for BCG, this implies that a 20-dose vial is not opened for just 1 child and BCG is often given only on specific days to assure that at least 10–12 children can be vaccinated [17]. Furthermore, some countries do not to give BCG at birth to LW infants. If BCG can substantially reduce neonatal mortality, it is important that WHO strongly encourage countries to follow their recommendations and provide BCG to all children immediately after birth, including LW children. Early administration of BCG should be part of the delivery services because that is where the first contact occurs for many children in low-income countries. Because boys have much higher mortality than girls in the first week of life, it is particularly important that boys receive BCG at birth. Receiving a BCG vaccine at birth might reduce the early mortality among boys to the same level as that of girls. Given the efficacy of BCG in reducing neonatal mortality it would be cost effective to open a vial of BCG even if only 1 child was present.

WHO and other policy makers should take this rapidly occurring effect of BCG into account when designing vaccination schedules. For example, if a new tuberculosis vaccine is introduced to replace BCG, the new vaccine should be evaluated not only for specific protection against tuberculosis but also for its nonspecific effect on neonatal mortality. The best option might well be to introduce a future tuberculosis vaccine for its effect against tuberculosis and maintain BCG at birth as an early “immune training” vaccine.

Supplementary Data

Supplementary materials are available at The Journal of Infectious Diseases online. Consisting of data provided by the authors to benefit the reader, the posted materials are not copyedited and are the sole responsibility of the authors, so questions or comments should be addressed to the corresponding author.

Notes

Transparency declaration. The first author affirms that the manuscript is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted and that any discrepancies from the study as planned have been explained.

Author contributions. P. A. and C. S. B. conceived and designed the BCG trials. S. B. S., I. M., and K. J. J. supervised the field data collection. S. B. S. analyzed the data. H. R. supervised the data analysis. S. B. S. wrote the first draft of the manuscript. All authors contributed to the final version of the manuscript.

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Potential conflicts of interest. All authors have completed the ICMJE uniform disclosure form at http://www.icmje.org/coi_disclosure.pdf and declare: no support from any
organization for the submitted work. The salaries for all authors are administered, but not financed, by SSI, the producer of the BCG vaccine SSI used in the present study; no other relationships or activities reported that could appear to have influenced the submitted work. All authors have submitted the ICMJE Form for Disclosure of Potential Conflicts of Interest. Conflicts that the editors consider relevant to the content of the manuscript have been disclosed.

References